## Chapter VI

## Parameters of a Random Variable

## D. 6. 1. (Expected Value)

The expected value of the random variable $X$, denoted by $E(X)$, is defined as

$$
E(X):= \begin{cases}\sum_{i=1}^{\infty} x_{i} \cdot p_{i} & \text { when } X \text { discrete } \\ \int_{-\infty}^{+\infty} x \cdot f(x) d x & \text { when } X \text { continuous }\end{cases}
$$

under the assumption

$$
\sum_{i=1}^{\infty}\left|x_{i}\right| \cdot p_{i}<\infty
$$

and

$$
\int_{-\infty}^{+\infty}|x| \cdot f(x) d x<\infty
$$

## R. 6. 1.

The expected value of a discrete random variable is the weighted mean of all outcomes $x_{i}$ of $X$ with the probabilities $p_{i}$ acting as weights.

## R. 6. 2.

The notion "expected value" in the probability theory has similarities with the notion "mean value", they are not, however, identical. The following example illustrates this fact:

## Ex. 6. 1.

A die will be tossed with the following outcomes:

$$
3,5,4,3,1
$$

The average mean is equal to

$$
\bar{x}=\frac{1}{5} \cdot(3+5+4+3++1)=3.2
$$

On the other hand, the expected value of the random variable

$$
X: \quad \text {,number of dots facing uppermost" }
$$

will be

$$
E(X)=\frac{1}{6} \cdot 1+\frac{1}{6} \cdot 2+\frac{1}{6} \cdot 3+\frac{1}{6} \cdot 4+\frac{1}{6} \cdot 5+\frac{1}{6} \cdot 6=3.5
$$

Whereas the mean value can vary from trial to trial, the expected value is an objective number independent of concrete outcomes of trials. The mean value approaches the expected value if the number of trials approaches infinity.

## Ex. 6. 2.

Given the probability density function

$$
f(x)=-\frac{2}{9} x+\frac{2}{3}, \quad x \in[0,3]
$$

find the expected value $E(X)$.

## Solution:

$$
\begin{aligned}
E(X) & =\int_{-\infty}^{+\infty} x \cdot f(x) d x=\int_{0}^{3} x \cdot\left(-\frac{2}{9} x+\frac{2}{3}\right) d x \\
& =\int_{0}^{3}\left(-\frac{2}{9} x^{2}+\frac{2}{3} x\right) d x \\
& =-\frac{2}{9} \cdot \frac{x^{3}}{3}+\left.\frac{2}{3} \cdot \frac{x^{2}}{2}\right|_{0} ^{3} \\
& =-\frac{2}{9} \cdot \frac{27}{3}+\frac{9}{3}=1
\end{aligned}
$$

## D. 6. 2. (Variance or Dispersion, Standard Deviation)

The dispersion or variance of the random variable $X$, denoted by $D^{2}(X)$, is defined as

$$
D^{2}(X):=E(X-E(X))^{2}
$$

i. e.

$$
D^{2}(X):= \begin{cases}\sum_{i=1}^{\infty}\left(x_{i}-E(X)\right)^{2} \cdot p_{i} & \text { when } X \text { discrete } \\ \int_{-\infty}^{+\infty}(x-E(X))^{2} \cdot f(x) d x & \text { when } X \text { continuous }\end{cases}
$$

under the assumption that the expected value exists.
The standard deviation, denoted by $D$, is defined as

$$
D(X):=\sqrt{D^{2}(X)} \quad(>0)
$$

## R. 6. 3.

The following relations can be easily verified:

$$
D^{2}(X)=E\left(X^{2}\right)-(E(X))^{2},
$$

i. e.

$$
D^{2}(X):= \begin{cases}\sum_{i=1}^{\infty} x_{i}^{2} \cdot p_{i}-\left(\sum_{i=1}^{\infty} x_{i} \cdot p_{i}\right)^{2} & \text { when } X \text { discrete } \\ \int_{-\infty}^{+\infty} x^{2} \cdot f(x) d x-\left(\int_{-\infty}^{+\infty} x \cdot f(x)\right)^{2} & \text { when } X \text { continuous }\end{cases}
$$

Ex. 6. 1. (continued)
Find the variance and the standard deviation of $X$.
Solution:

$$
D^{2}(X)=(1-3.5)^{2} \cdot \frac{1}{6}+(2-3.5)^{2} \cdot \frac{1}{6}+\ldots+(6-3.5)^{2} \cdot \frac{1}{6} \approx 2.92
$$

or

$$
\begin{aligned}
& D^{2}(X)=1 \cdot \frac{1}{6}+4 \cdot \frac{1}{6}+\ldots+36 \cdot \frac{1}{6}-3.5^{2} \approx 2.92 . \\
& D(X) \approx 1.71
\end{aligned}
$$

Ex. 6. 2. (continued)
Find the variance and the standard deviation of $X$.
Solution:

$$
\begin{aligned}
D^{2}(X) & =\int_{0}^{3}(x-1)^{2} \cdot\left(\frac{2}{9} x+\frac{2}{3}\right) d x \\
& =\int_{0}^{3}\left(-\frac{2}{9} x^{3}+\frac{10}{9} x^{2}-\frac{14}{9} x+\frac{2}{3}\right) d x \\
& =-\frac{2}{9} \cdot \frac{x^{4}}{4}+\frac{10}{9} \cdot \frac{x^{3}}{3}-\frac{14}{9} \cdot \frac{x^{2}}{2}+\left.\frac{2}{3} x\right|_{0} ^{3} \\
& =\frac{1}{2}
\end{aligned}
$$

or

$$
\begin{aligned}
D^{2}(X) & =\int_{0}^{3} x^{2} \cdot\left(-\frac{2}{9} x+\frac{2}{3}\right) d x-1 \\
& =\int_{0}^{3}\left(-\frac{2}{9} x^{3}+\frac{2}{3} x^{2}\right) d x-1 \\
& =-\frac{2}{9} \cdot \frac{x^{4}}{4}+\left.\frac{2}{3} \cdot \frac{x^{3}}{3}\right|_{0} ^{3}-1 \\
& =\frac{1}{2}
\end{aligned}
$$

## R. 6. 4.

Let $X$ and $Z$ be two random variables. Given the transformation

$$
Z=g(X)
$$

we would like to derive informations about $Z$ assuming we know the distribution of $X$. This will be illustrated by the following example:

## Ex. 6. 3.

Let us consider the function

$$
Z=4 X
$$

with the following probability and distribution functions for the random variable $X$ :

| $x_{i}$ | 2 | 4 |
| :---: | :---: | :---: |
| $P\left(X=x_{i}\right)$ | $\frac{1}{4}$ | $\frac{3}{4}$ |

$$
F(x):=\left\{\begin{array}{lll}
0 & \text { when } & x \leq 2 \\
\frac{1}{4} & \text { when } & 2<x \leq 4 . \\
1 & \text { when } & x>4
\end{array}\right.
$$

We are interested in finding the probability and distribution functions for the random variable $Z$ :

$$
\begin{aligned}
P\left(X=x_{i}\right) & =P\left(4 X=4 x_{i}\right) \\
& =P\left(Z=4 x_{i}\right) \\
& =P\left(Z=z_{i}\right), \quad i=1,2,
\end{aligned}
$$

and

$$
\begin{aligned}
F(z) & =P(Z<z) \\
& =P(4 X<z) \\
& =P\left(X<\frac{z}{4}\right) \\
& =P\left(\frac{z}{4}\right)=F(x) .
\end{aligned}
$$

Thus, we have the probability function:

| $z_{i}$ | 8 | 16 |
| :---: | :---: | :---: |
| $P\left(Z=z_{i}\right)$ | $\frac{1}{4}$ | $\frac{3}{4}$ |

and the distribution function:

$$
F(z):=\left\{\begin{array}{lll}
0 & \text { when } & z \leq 8 \\
\frac{1}{4} & \text { when } & 8<z \leq 16 \\
1 & \text { when } & z>16
\end{array} .\right.
$$

Let us now calculate the expected value of $Z$ :

$$
\begin{aligned}
E(Z) & =z_{1} \cdot P\left(Z=z_{1}\right)+z_{2} \cdot P\left(Z=z_{2}\right) \\
& =8 \cdot P(Z=8)+16 \cdot P(Z=16) \\
& =8 \cdot P(4 X=8)+16 \cdot P(4 X=16) \\
& =8 \cdot P(X=2)+16 \cdot P(X=4) \\
& =8 \cdot \frac{1}{4}+16 \cdot \frac{3}{4} \\
& =14 .
\end{aligned}
$$

## R. 6. 5.

Generalising the results of the above example and assuming that the corresponding expected values exist, the following relations can be proved:

$$
E(Z)=E(g(x))= \begin{cases}\sum_{i=1}^{\infty} g\left(x_{i}\right) \cdot P\left(X=x_{i}\right) & \text { when } X \text { discrete } \\ \int_{-\infty}^{+\infty} g(x) \cdot f(x) d x & \text { when } X \text { continuous }\end{cases}
$$

## Ex. 6. 4.

Consider the random variable $X$ with the probability function

| $x_{i}$ | $x_{1}$ | $x_{2}$ | $\cdots$ | $x_{n}$ |
| :---: | :---: | :---: | :---: | :---: |
| $p_{i}=P\left(X=x_{i}\right)=f\left(x_{i}\right)$ | $p_{1}$ | $p_{2}$ | $\cdots$ | $p_{n}$ |

Let

$$
Z=a X+b, \quad a, b=\text { const } .
$$

Then, we have:

$$
\begin{aligned}
E(Z) & =\sum_{i=1}^{\infty}\left(a x_{i}+b\right) \cdot p_{i} \\
& =a \cdot \sum_{i=1}^{\infty} x_{i} \cdot p_{i}+b \cdot \sum_{i=1}^{\infty} p_{i} \\
& =a \cdot \sum_{i=1}^{\infty} x_{i} \cdot p_{i}+b .
\end{aligned}
$$

Hence, assuming that $E(X)$ exists, we have

$$
E(a X+b)=a \cdot E(X)+b .
$$

Similarly, it can be proved:

$$
D^{2}(a X+b)=a^{2} \cdot D^{2}(X)
$$

assuming that $D^{2}(X)$ exists.

## Ex. 6. 5.

Consider the random variable $X$ with

$$
E(X)=\mu, \quad D^{2}(X)=\sigma^{2} \quad\left(\sigma^{2} \neq 0\right)
$$

For the function

$$
Z=g(X):=\frac{X-\mu}{\sigma}=\frac{1}{\sigma} \cdot X-\frac{\mu}{\sigma}
$$

we obtain

$$
\begin{aligned}
E(Z) & =\frac{1}{\sigma} \cdot E(X)-\frac{\mu}{\sigma} \\
& =\frac{1}{\sigma} \cdot \mu-\frac{\mu}{\sigma}=0 \\
& =0
\end{aligned}
$$

$$
\begin{aligned}
D^{2}(Y) & =\frac{1}{\sigma^{2}} \cdot D^{2}(X) \\
& =\frac{\sigma^{2}(X)}{\sigma^{2}(X)} \\
& =1
\end{aligned}
$$

## D. 6. 3. (Standardisation, Standardised Random Variable)

The random variable $Z$ is called a standardised random variable if

$$
E(Z)=0, \quad D^{2}(Z)=1 .
$$

The process

$$
Z=g(X):=\frac{X-\mu}{\sigma}=\frac{1}{\sigma} \cdot X-\frac{\mu}{\sigma}
$$

is called standardisation.
(Last revised: 03.06.08)

